

Validation Procedures in Radiologic Diagnostic Models

Neural Network and Logistic Regression

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Arana E, Delicado P, Martí-Bonmatí L. Validation procedures in radiologic diagnostic models: Neural network and logistic regression. Invest Radiol 1999;34:636-642.

OBJECTIVE. To compare the performance of two predictive radiologic models, logistic regression (LR) and neural network (NN), with five different resampling methods.

METHODS. One hundred sixty-seven patients with proven calvarial lesions as the only known disease were enrolled. Clinical and CT data were used for LR and NN models. Both models were developed with cross-validation, leave-one-out, and three different bootstrap algorithms. The final results of each model were compared with error rate and the area under receiver operating characteristic curves (A_z).

RESULTS. The NN obtained statistically higher A_z values than LR with cross-validation. The remaining resampling validation methods did not reveal statistically significant differences between LR and NN rules.

CONCLUSIONS. The NN classifier performs better than the one based on LR. This advantage is well detected by three-fold cross-validation but remains unnoticed when leave-one-out or bootstrap algorithms are used.

KEY WORDS. Skull, neoplasms; statistics, logistic regression; neural networks; receiver operating characteristic curve; statistics, resampling.

PREDICTIVE MODELS have been extensively studied as supportive diagnostic aids to radiology in a variety of diseases.¹ Among these methods, neural networks (NNs) have been developed for radiologic purposes. They are based on a parallel architecture with several layers. Each layer receives input only from the directly preceding one, and adjacent layers are fully connected. The utility and flexibility of NNs arise from the application of learning algorithms that allow the network to construct the correct weights and, hence, the desired function for a given set of observations. Although several algorithms have been discussed in the literature, the most commonly employed is the back-propagation of errors. Logistic regression (LR) is a nonlinear regression technique that has proven to be very robust in a number of medical domains and is acknowledged as the statistical analysis of choice for predicting dichotomous outcomes.²

A standard procedure for evaluating the performance of a model would be to split the data into a training set, a cross-validation set (used to determine the stopping point to avoid overfitting, or used to set additional parameters, such as weight elimination), and a test set. The test set is a set of examples, not previously shown to the NN, used only to assess the performance (generalization) of a fully specified classifier. Regression models should generally use a training and a test set as well.³ In medical settings, often because of the small amount of data available for training these models, new cases are rarely found to be tested. For these reasons, validation procedures with resampling techniques are usually employed. These techniques use part of the data set to train and to validate these models. Recent radiologic papers dealing with NNs have discussed these issues. However, many researchers are using such models without validating the necessary assumptions.⁴ Models developed this way are

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The authors thank John Boone, PhD, Professor of Radiology, University of California Davis Medical Center, for his critical review of and suggestions for this manuscript, and the two reviewers for their ideas to improve our work.

A portion of this paper represents the first author's doctoral thesis research.

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Received February 3, 1999, and accepted for publication, after revision, June 1, 1999.

TABLE 1. Variables Recorded and Their Description

Variables		Values
1	Age	Years
2	Gender	Male, female
3	First symptom noticed	Tumour, pain, headache, incidental finding, others
4	Symptoms length	Months
5	Number of lesions	Number
6	Bone	Frontal, parietal, occipital, sutures or fontanels
7	Centricity	Outer table, diploe, inner table, intracranial, extracranial
8	Maximal diameter	In millimeters
9	Shape	Circular, ovoid (image in plane of greatest diameter)
10	Character-appearance	Lytic permeative, lytic moth-eaten, lytic geographic, blastic, Mixed blastic permeative, mixed blastic moth-eaten, mixed blastic geograph
11	Expansivity	No, mild, moderate, severe
12	Edge definition	Poor, moderate, well
13	Lobularity	Lobular, smooth
14	Marginal sclerosis	No, partial, rind (<2 mm), band (>2 mm), no applicable
15	Periosteal reaction	No, yes (any form)
16	Matrix	None, ground glass, calcified, ossified and sequestration
17	Cortical involvement	Diploe, internal, external, both corticals
18	Form of cortical involvement	None, thickened, thinned, broken
19	Swell/mass	None, intracranial, subgaleal, both

unlikely to stand the test validation on a separate patient sample.³ The main NN drawbacks are that usually they are not matched with statistical techniques of reference, and they are used in small samples.² The associated risk of overfitting on noisy data is a major concern in NN design,⁵ although LR methods can suffer from the same problems.²

Calvarial lesions are often found during CT imaging of the brain with no specific symptoms. The signs for their characterization are based on imaging studies, especially CT. A diagnostic model could help the radiologist with these uncommon lesions.⁶ The purpose of this work was to study the ability of five resampling methods, leave-one-out, and bootstrap, for validating LR and NN models to classify calvarial lesions and compare their results.

Materials and Methods

Population

A complete discussion of the population and methodology has been previously presented.⁶ As part of our review, 167 patients with calvarial focal bone lesions were reviewed for a 4-year period from 4012 head CT scans. The lesions were analyzed in detail and reviewed by two of us by consensus (EA and LM-B). All patients were examined with at least two plain radiographic projections and CT. CT sections were obtained with the window width and level settings that best allowed evaluation of soft tissue structures and bones. The setting on each scanner was individualized for every patient.

There were 74 male patients (44.6%) and 93 female patients (55.4%) with an age range of 0.5 to 81 years (31.1 ± 25.5 years, mean $\pm$ SD). The total number of benign lesions was 122 (73.1%), with an age range from 5 to 80 years (26.1 ± 22.9). There were 45 malignant ones

(27.0%), with an age range from 0.5 to 84 years (55.7 ± 18.6).

Explanatory Diagnostic Variables

Nineteen morphologic imaging characteristics as well as anatomic and demographic data were evaluated without knowledge of the final diagnosis (Table 1). All findings were recorded for all patients in a spreadsheet and used for both the LR model and the NN analysis. When a single feature had two or more findings in the same lesion, the most severe form was the one recorded. There were no missing data. Lesions were divided into benign and malignant. Of the 19 explanatory variables, 3 were continuous,^{1,4,8} 1 was quantitative discrete,⁵ 2 were ordinal,^{11,12} and the rest were qualitative. For the latter, we used a 1-out-of-C code, where a variable with C categories is converted into C Boolean inputs, each of which is high for a certain category; eventually 43 explanatory variables were used.

Logistic Regression

The LR equation assumes that the expected probability of a dichotomous outcome is as follows:

$$P = \frac{1}{1 + e^{-(\beta_0 + \beta_1 X_1 + \beta_2 X_2 + \dots + \beta_p X_p)}}$$

where the X_i are variables with numeric values (if dichotomous, they are, for example, 0 for false and 1 for true) and the β_i are the regression coefficients that quantify their contribution to the probability. LR has proven to be an effective way of estimating probabilities for dichotomous variables, in this case the probability of benign or malignant.

We fit the LR model by a Newton-Raphson type of

iterative algorithm implemented in MATLAB 4.0 (MathWorks, Inc., Natick, MA). The 19 explanatory variables listed in Table 1 were recodified as 43 input variables, and all of them were used in the β_i parameters fitting. The estimation procedure also provides a measurement of the estimated coefficients' standard errors. This allows defining confidence intervals (CI) for the coefficients and quantifying whether they are significantly different from zero. The value of a standardized coefficient is an indication of the relevance of the corresponding explanatory variable. The usual output of an LR estimation is the odd-ratios list (ie, the values e^{β_i}) and the CI for them, derived from the coefficients' CI (also by using the exponential function).

Neural Network

The developed NN had three layers, with a feed-forward architecture, being trained by the back-propagation algorithm with the sigmoid activation function.⁷ The structure of the NN included 43 input units, which were the same as in the LR model. No well-established theoretical method exists for designing an ideal NN, and the optimal number of hidden nodes and iterations is unknown.^{7,8} The number of hidden neurons was chosen from those with a satisfactory adjustment of the NN. The number of tested hidden neurons ranged from 10 to 30, and a satisfactory trade-off between low miss-classification rate and complex design was obtained with 15 hidden neurons. Two stopping rules were implemented, one based on the number of iterations (the maximum allowed number was 500) and the other based on a lower bound for the mean squared error that indicates how well the outputs are calibrated to their targets (this bound was equal to 0.03). In fact, the effective stopping rule was always the first one in all the training procedures run. The main drawback of these stopping criteria is the possibility of finding a local optimum instead of the global optimal value. This is a common problem for a wide variety of multivariate optimization procedures. Some techniques including random perturbations of intermediate solutions (as simulated annealing⁹) are specially designed to lead to a global optimum. We limited ourselves to standard implemented training processes in the NN MATLAB toolbox.

The main concern in a design with an excessive number of hidden nodes is overfitting. The number of hidden nodes determines the complexity of the functions represented by an NN: as this number increases, the function is more complex. If a net contains too many hidden nodes, this net can learn the training set so perfectly that even a zero training error could be achieved (then we say that the net overfits the training set); however, this net will usually have a large generalization error in independent test sets. This situation will appear because, at least in theory, we take the best function among too many flexible classes of functions. In practice, only one function is taken when an NN is trained: the optimization procedure is limited by the specific

training algorithm and stopping rules implemented (eg, number of iterations, required level of precision). Therefore, not only the number of hidden nodes, but also the training algorithm, determines the complexity of the NN. In our case, we use only 500 iterations in the training procedure. Because not every function with 15 nodes can be learned in 500 iterations, the class of actual possibly learned functions is not so large; as a consequence, overfitting problems may not appear.

Comparison and Performance

One of the performance measures of a classification rule is the probability of miss-classifying a new observation, assuming that a case is assigned to a class if the classification rule gives it a probability >0.5 to belong to that class. A naive estimator of this probability is the apparent classification error or error rate, defined as the number of incorrectly classified cases in both classes divided by the number of total cases. This estimator is optimistically biased because the same cases are used to fit the classifier and to compute the error rate.

Another measure of performance is the area (A_z) under the receiver operating characteristic (ROC) curve. The continuum output given by LR and NN was compared with their correct values to obtain the ROC curves. ROC curves measure predictive utility by showing the trade-off between the true-positive rate (sensitivity, probability of correct classification of a positive case) and the false-positive rate ($1 - \text{specificity}$, probability of incorrect classification of a negative case) inherent in selecting specific thresholds on which predictions might be based. The area under this curve represents the probability that, given a positive case and a negative one, the classifier rule output will be higher for the positive case, and it is not dependent on the choice of decision threshold. This way is less dependent on the frequency of malignancy in the population and allows consideration of the sensitivity and specificity of the model at various probability levels. In this study, the area under the ROC curve was obtained by plotting sensitivity versus $1 - \text{specificity}$ for each possible predictive score cutpoint and summing the areas of the created trapezoids. Statistical differences and confidence intervals between the NN and LR outputs were compared with a two-tailed, nonparametric approach according to the method described by Hanley and McNeil.^{10,11}

Validation

Pure naive validation methods use all the cases to build the models and also to validate them. Because the estimated model has been fitted to the idiosyncrasies of the training sample, the validation based on the same sample tends to underestimate the probability of misclassification. To estimate the performance of a classifier, a validation method with lower bias is preferred. For these reasons, resampling methods are used to depict its future prediction accuracy

and also to choose a classifier from a given set (model selection), or to combine classifiers.

Cross-Validation With Random Subsampling

In k -fold cross-validation, the data set is randomly split into k mutually exclusive subsets (the folds) $D_1, D_2, \dots, D_k$ of approximately equal size. The classifier rule is trained and tested k times. The cross-validation estimation of accuracy is the overall number of correct classifications, divided by the number of instances in the data set. Because we have a moderate-size sample ($n = 167$), we developed a cross-validation with $k = 3$. Every time, a model is developed with $n_1 = 112$ and tested in the rest $n_2 = 55$. This way, we are in the ideal situation of having an independent sample to test the model. This process is repeated 20 times, with randomly chosen training and testing sets producing 60 unbiased estimates of discriminant ability. Because the test samples are independent of the training data, the results derived from this three-fold cross-validation are reliable. The drawback for obtaining this reliability is that a third of the data in the model estimation phase is lost. Error rates are the average of the 20 resampling process.

Leave-One-Out

Several validation methods are available if one cannot afford to lose a significant part of the sample in the estimation step. One of them is leave-one-out, and others are based on bootstrap principles. In a sample of size n , leave-one-out is n -fold cross-validation. According to this method, the entire database ($n = 167$) but one patient is used to develop the classifying models. Then, the LR or NN model is tested on the patient left out. The same process is repeated so that each pattern of the data is left out once.

Bootstrap

According to the bootstrap method, the training set (a bootstrap sample) is generated by sampling with replacement n times from the available n cases. A diagnostic model is trained on the bootstrap model and then tested on both the bootstrap and the original set, and accuracy is twice measured. The difference between both rates of misclassifications reflects the optimism of the naive apparent error rate. The same process is repeated B times, and the average of those differences is taken as a global measure of the optimism. The estimates developed also include two different bootstrap algorithms¹²:

Apparent error rate + optimism (bootstrap 1)

Error rate 0.632 bootstrap (bootstrap 2)

Error rate 0.632 + bootstrap. This combines the leave-one-out bootstrap with a measure of over fitting (bootstrap 3).

The number of B subsamples generated was 100, and results were obtained with the Bootpred routine written in S-plus and described by Efron and Tibshirani.¹³

TABLE 2. Most Relevant Variables Included for the Logistic Regression Model ($P < 0.05$)

Variable	Odds ratio	CI 95%
Age	1,117	1,036–1,203
Character-appearance		
Mix. blastic permeative	0,004	0,000–0,903
Periosteal reaction	0,041	0,001–1,250
Symptoms length	0,949	0,894–1,006
First symptom noticed		
Tumour	34,917	0,495–2464,392
Character-appearance		
Mix. blastic moth-eaten	0,009	0,000–2,770
Centricity		
Outer table	16,555	0,544–503,762

CI: confidence interval.

Results

Logistic Regression and Neural Network Fit

Only some numeric results from the estimation phase are reported here, because our main objectives were the comparison among different resampling procedures. Table 2 shows the most relevant variables in the logistic model fit, sorted by the standardized coefficient values. Odds ratios and 95% CI for them are shown. Only the first two variables (age and mixed blastic permeative character appearance) have CI for odds ratios without including the value 1. The results obtained in the NN and LR fitting, which are listed later, allowed us to state that the designed NN (with 15 hidden nodes and 500 iterations) does not overfit the data because LR presents a lower training error than NN fit, but LR has a greater generalization error.

Pure Naive Models

The apparent error rate for the LR model was 0.0240, with an A_z of 0.9993 ± 0.0036 . The NN showed an error rate of 0.0599 with an A_z of 0.9505 ± 0.0285 . These results reveal the overfitting of the models without resampling methods.

Cross-Validation with Random Subsampling

ROC A_z with their CIs for the different resampling methods are shown in Fig. 1 and Table 3. Both models, particularly NN, obtained the smallest variances among the other resampling methods (Fig. 1). The logit model performed worse than NN, with higher error rates of 0.1916 versus 0.1377 ($P < 0.0001$), and A_z 0.8103 versus 0.8854 ($P < 0.001$) (Fig. 2). The probability values reveal larger differences when error rates are compared instead of A_z .

Leave-One-Out

The LR using the leave-one-out method performed significantly better than the cross-validation in terms of error rate ($P < 0.01$). Although the NN obtained a higher A_z and lower overall error rates than the logit model, there were no statistical differences (Table 3). Areas under ROC curves presented

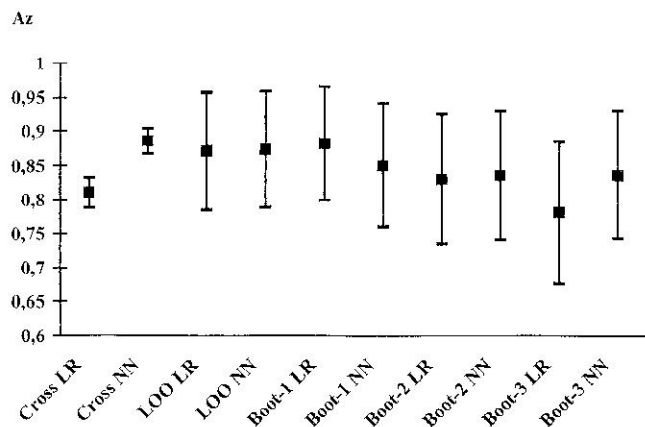


Figure 1. ROC Az with their confidence interval (95%) for the different resampling methods. Cross: cross-validation; LOO: leave-one-out; Boot-1: bootstrap-1; Boot-2: bootstrap-2; Boot-3, bootstrap-3.

markedly larger variances than with cross-validation, although smaller than those seen with bootstrap (Fig. 1).

Bootstrap

There were no statistical differences between the LR and NN models validated with the three bootstrap algorithms either in error rates or in A_z . The third bootstrap approach (0.632+bootstrap) gave similar results to three-fold cross-validation (Fig. 3). The opposite happened when the results from bootstrap 1 were examined (LR performed marginally better than NN), and bootstrap 2 indicates advantages for NN in terms of the global measure A_z but not in terms of error rate (Table 3). Variance of A_z estimations was progressively greater in every bootstrap algorithm, showing profound differences with cross-validation algorithms (Fig. 1).

Discussion

There is a great interest in the performance of NNs and classifier rules in medical applications. Ideally, we would train the NN and LR models with a larger data set and apply the trained models to an entirely different data set to evaluate its performance. However, we were not able to split our database into completely separate training and testing data

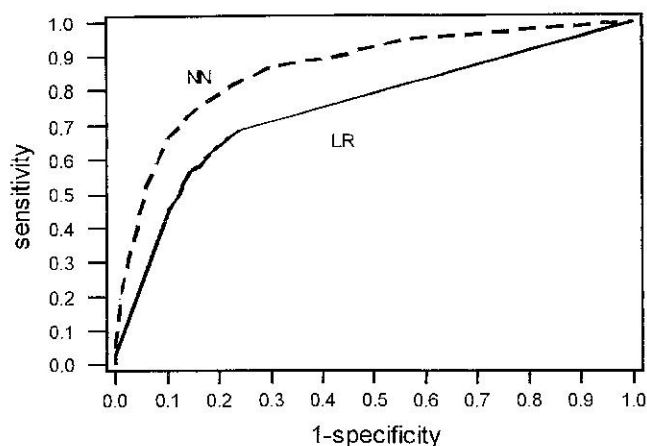


Figure 2. ROC curves for LR and NN based on threefold cross-validation. The reported curves are the average of the curves obtained in the twenty resampling processes.

sets because the number of confirmed lesions is currently small (but to our knowledge the largest series in the literature). Future research is needed to compare trained models on completely separate data sets. Although some theories have been presented on sampling methods, these are still in their infancy.¹⁴

The objectives of these studies on predictive models should be clarified, because the approaches are different if we want to find the "gold standard predictive" model or a model whose prediction error is the lowest. The classification performance of stochastic models, such as LR and NN, however, depends on the estimation techniques. Thus, error rate is optimistically biased and A_z is not. This may be responsible for the greater differences between LR and NN when they are compared in terms of error rates.

In the medical practice, data set size is always finite and usually smaller than desired. The main drawback of k -fold cross-validation is that it makes inefficient use of the data: a third of the data set is not used to train the classifier rule. Previous studies on this method have shown that as the training sample size increased, so did the predictive accuracy of the NN.¹⁵ One may expect that these error rates would decrease and, respectively, A_z would increase, when

TABLE 3. Results of the Logistic Regression and Neural Network Models With the Different Resampling Methods in Terms of Overall Error Rate and Area Under the ROC Curve (A_z)

Algorithm	Logistic regression		Neural network	
	Error rate (CI 95%)	ROC Az (CI 95%)	Error rate (CI 95%)	ROC Az (CI 95%)
Cross validation	0.1916 (0.1810, 0.2080)	0.8103 (0.7883, 0.8323)	0.1377 (0.1240, 0.1476)	0.8854 (0.8674, 0.9034)
Leave one out	0.1377 (0.0854, 0.1899)	0.8711 (0.7854, 0.9567)	0.1198 (0.0705, 0.1690)	0.8736 (0.7887, 0.9584)
Bootstrap-1	0.0958 (0.0539, 0.1446)	0.8819 (0.7994, 0.9644)	0.1737 (0.1162, 0.2311)	0.8508 (0.7600, 0.9415)
Bootstrap-2	0.1377 (0.0840, 0.1880)	0.8303 (0.7350, 0.9256)	0.1676 (0.1128, 0.2267)	0.8351 (0.7408, 0.9294)
Bootstrap-3	0.1736 (0.1169, 0.2321)	0.7809 (0.6766, 0.8852)	0.1676 (0.1128, 0.2267)	0.8361 (0.7420, 0.9302)

CI: confidence interval.

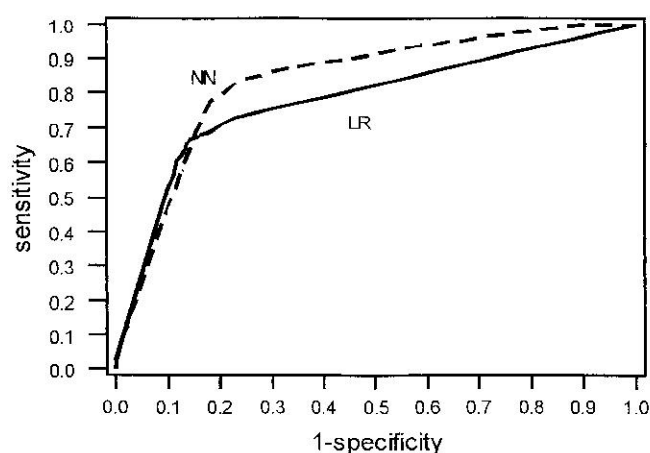


Figure 3. ROC curves for LR and NN based on .632+ bootstrap.

the whole sample is used to estimate both LR and NN. To extrapolate these results from our sample size of 112 to 167, the relevance of the numeric values is more qualitative (NN clearly surpassed LR) than quantitative. Therefore, the error rates obtained may be far from the reliability with larger samples.

According to our results, leave-one-out has been described as an almost unbiased method but one that has high variance.¹⁶ Leave-one-out produced similar variance to the other resampling methods except for cross-validation. The leave-one-out procedure ensures that regardless of the sample size, the relevant observation would not be in the pseudooptimal solutions more than once.¹⁴ This resampling technique can easily produce significantly different results in NN settings, depending on the training-stopping criterion.¹⁵ A theoretical study suggested that cross-validation and leave-one-out do not offer significant improvement over the apparent error, whereas the improvement given by bootstrap is substantial.¹⁷ Nevertheless, these comparisons were carried out using simulated data and the root mean squared error for performance measuring, instead of real data and ROC analysis. Cross-validation can be very sensitive to the specific sample splitting. Further, if the specific test set is given and the data are sparse and noisy (as in medical settings), tests of predictive reliability may not present a good picture of sample variability, or potential changes in model specification.¹⁸

NNs validated with cross-validation in radiologic diagnoses have shown protean features. Theory has been corroborated with real data—as the training sample size increased, so did the network's predictive accuracy, for instance in ventilation-perfusion imaging.^{15,19} However, in other fields, such as focal bone lesions, the performance appeared to be more strongly related to how radiographically distinctive each pathologic type was rather than the number of cases available.²⁰ Therefore, simulation studies concluded that cross-validation gives a nearly unbiased es-

timate of error, but often with unacceptably high variability, particularly if the database is small.²¹ In contrast, our results showed that cross-validation provided the smallest variance among the different algorithms. The reason for our small variance compared with those of Tourassi and Floyd¹⁵ and Efron²¹ is that we randomly repeated 20 times the three-fold cross-validation procedure, obtaining 60 observations of cross-validation error to be averaged, far more than they did. Efron suggested that different resampling methods applied to practical situations could provide different answers. However, the results were very similar among more recent papers comparing resampling methods on radiologic data (Tourassi and Floyd and this work).

Bootstrap techniques in radiologic diagnosis have been described in the report by Tourassi and Floyd.¹⁵ In 1997, Efron and Tibshirani¹² proposed the “.632+” estimator, which combines the leave-one-out bootstrap with a measure of overfitting. In extensive simulations, it has been the best-performing bootstrap and offer some advantages over cross-validation. Similar results appear in our study: “.632+” was the only bootstrap-based procedure that produced results similar to those obtained by three-fold cross-validation. The available asymptotic results show its validity for a large number of linear, nonlinear, and even nonparametric regression problems. In contrast to the bayesian approach, no distributional assumptions (eg, normal errors) have to be specified. For a large sample size and a small B value, bootstrap does not ensure that all the relevant observations will be deleted at least once. The advantage is that in small-size samples and with a relatively large B value, bootstrap algorithms may capture the variability of sample data better than the leave-one-out procedure.¹⁴ Particularly in NN settings, bootstrap yields rather reliable estimates of the variance even in small-sample situations, models where the distribution of residual depends on the input; in addition, it is more robust if the selected model is incorrect.¹³ Therefore, it does not need to assume normality or symmetry in the data.

A more practical question that should be considered as well is whether the bootstrap is worth the extra computer time required. Each bootstrap iteration requires a run of the algorithm, and it seems unlikely that this can be improved on. In our work, all bootstrap routines (with B = 100) were the second-longest after leave-one-out (n = 167) with the NN (overall duration, respectively, 115'41" and 180'10"). This is just a manifestation of Occam's razor, which states that complex models should not be preferred over simpler ones.²²

We agree with a previous paper dealing with the uncertainty of choosing resampling methods on NN design.¹⁵ This paper stated that we can be confident about the classifier rule performance if the estimates of resampling methods are very similar. However, other types of NNs may show different levels of performance due to the many pa-

rameters involved in NN development (ie, the leave-one-out and the training-stopping criteria). Training problems are common and may be difficult to address even with state-of-the-art resampling methods. The required number of training cases clearly depends on the difficulty of the decision task, the amount of input, and the number of hidden and output nodes (and also the number of weights). Thus, the information provided to the network is crucial. Reinus et al²⁰ have shown improved performance of their NN design using a greater number of input units. This means that the imaging parameters were more explicitly defined than in other NNs with fewer input features.

Nevertheless, one work that evaluated forecasts on financial data found that the variation from different resampling (ie, splits between training, cross-validation, and test sets) is significantly larger than the variation from different network conditions (such as architecture and initial weights).¹⁸

In summary, NN classification rule is preferred to LR in the diagnosis of focal calvarial lesions. This advantage is well detected by three-fold cross-validation but remains unnoticed when leave-one-out or bootstrap algorithms are used. However, both leave-one-out and "0.632+" bootstrap slightly indicate the superiority of NN over LR.

References

1. Wu Y, Giger ML, Doi K, Vybomy CJ, Schmidt RA, Metz CE. Artificial neural networks in mammography: Application to decision making in the diagnosis of breast cancer. *Radiology* 1993;187:81-87.
2. Tu J. Advantages and disadvantages of using artificial neural networks versus logistic regression for predicting medical outcomes. *J Clin Epidemiol* 1996;49:1225-1331.
3. Harrell F, Lee K, Matchar D, Reichert T. Regression models for prognostic prediction: Advantages, problems and suggested solutions. *Cancer Treat Rep* 1985;69:1071-1077.
4. Christy P, Tervonen O, Scheithauer B, Forbes G. Use of neural network and a multiple regression model to predict histologic grade of astrocytoma from MRI appearances. *Neuroradiology* 1995;37:89-93.
5. Geman S, Bienenstock E, Doursat R. Neural networks and the bias/variance dilemma. *Neural Computation* 1992;4:1-58.
6. Arana E, Marti-Bonmati L, Paredes R, Bautista D. Focal calvarial bone lesions. Comparison of logistic regression and neural network. *Invest Radiol* 1998;33:738-745.
7. Penny W, Frost D. Neural networks in clinical medicine. *Med Decis Making* 1996;16:386-398.
8. Miller A, Blott B. Review of neural network application in medical imaging and signal processing. *Med Biol Eng Comput* 1992;30:449-464.
9. Ripley BD. Pattern recognition and neural networks. Cambridge University Press, 1996.
10. Hanley J, McNeil B. The meaning and use of the area under a receiver operating characteristic (ROC) curve. *Radiology* 1982;143:29-36.
11. Hanley J, McNeil B. A method of comparing the areas under receiver operating characteristic curves derived from the same cases. *Radiology* 1983;148:839-843.
12. Efron B, Tibshirani R. Cross-validation and the bootstrap: Estimating the error rate of a prediction rule. *J Am Stat Assoc* 1997;92:548-560.
13. Efron B, Tibshirani R. Regression models. In: Efron B, Tibshirani R, eds. *An introduction to the bootstrap*, 1st ed. London: Chapman & Hall, 1993:105-123.
14. Ziari H, Leatham D, Ellinger P. Development of statistical discriminant mathematical programming model via resampling estimation techniques. *Am J Agr Econ* 1997;79:1352-1362.
15. Tourassi GD, Floyd CE. The effect of data sampling on the performance evaluation of artificial neural networks in medical diagnosis. *Med Decis Making* 1997;17:186-192.
16. Kohavi R. A study of cross-validation and bootstrap for accuracy estimation and model selection. Technical report. Computer Science Department, Stanford University, 1995.
17. Gong G. Cross-validation, the jackknife, and the bootstrap: excess error estimation in forward logistic regression. *J Am Stat Assoc* 1986;81:108-113.
18. LeBaron B, Weigend AS. Evaluating neural network predictors by bootstrapping. Technical report, University of Wisconsin, 1995.
19. Fisher R, Scott J, Palmer E. Neural networks in ventilation-perfusion imaging. Part I. Effects of interpretative criteria and network architecture. *Radiology* 1996;198:699-706.
20. Reinus WR, Wilson AJ, Kalman B, Kwasny S. Diagnosis of focal bone lesions using neural networks. *Invest Radiol* 1994;29:606-611.
21. Efron B. Estimating the error rate of a prediction rule: Improvement on cross-validation. *J Am Stat Assoc* 1983;78:316-330.
22. MacKay DJC. Bayesian model comparison and backprop nets. In: Moody J, Hanson S, Lippman R, eds. *Advances in neural information processing systems 4*. San Mateo: Morgan Kaufmann, 1992:839-846.